



Economic Growth and Income Inequality in Sub-Saharan Africa: Does Environmental Policy Matter?

Pilanyomon Richard COULIBALY ^a
and Wadjamsse Beaudelaire DJEZOU ^{a*}

^a *Laboratory of Analysis and Economics Policy Modelling (LAMPE), Alassane Ouattara University, Côte d'Ivoire.*

Authors' contributions

This work was carried out in collaboration between both authors. Both authors read and approved the final manuscript.

Article Information

DOI: <https://doi.org/10.9734/jemt/2024/v30i111253>

Open Peer Review History:

This journal follows the Advanced Open Peer Review policy. Identity of the Reviewers, Editor(s) and additional Reviewers, peer review comments, different versions of the manuscript, comments of the editors, etc are available here: <https://www.sdiarticle5.com/review-history/127028>

Original Research Article

Received: 10/09/2024

Accepted: 14/11/2024

Published: 19/11/2024

ABSTRACT

The counter-intuitive correlation between economic growth and income inequality in sub-Saharan Africa in recent years and the lack of consensus in the literature on this relationship call for its re-examination, particularly in the context of sustainable development. This article analyses the effect of environmental policies on the relationship between economic growth and income inequality in sub-Saharan Africa. With its suitability for analyzing non-linear relationships, the panel smooth transition regression model (PSTR) of Gonzalès (2005) was applied to data covering the period 1999 to 2021 from the World Bank, the Yale Center for Environmental Law and Policy (YCELP), the World Governance Indicator (WGI) and the World Income Inequality Database (WIID). In the presence of a strict application of environmental policies, the results show a U-inverted relationship

*Corresponding author: E-mail: wjdzou@gmail.com;

between economic growth and income inequality in sub-Saharan Africa. The more a relatively wealthy country adopts a rigorous environmental policy, the more economic growth contributes to reduce income inequalities. Thus, the socially virtuous effect of a rigorous environmental policy only occurs in economically dynamic countries. In other words, environmental policies in sub-Saharan African countries can only be promoted in relation to their economic performance. Any standardisation of environmental policies on a continental or sub-regional scale for economically different countries should be reconsidered.

Keywords: Environmental policy; economic growth; income inequality; sub-Saharan Africa; PSTR.

JEL Code: N34, Q51, B22

1. INTRODUCTION

Economic growth and income inequality are concerns for decision-makers in both developed and developing countries. Despite strong economic growth in sub-Saharan Africa in recent years, income inequalities have remained stable in some countries and increased in others. Indeed, despite an economic growth of about 3 percentage points (1.3% to 4%) between 2016 and 2021 (World Bank, 2022), sub-Saharan Africa remains one of the regions most affected by income inequality, with a Gini index of 0.46 compared to 0.48 at the global level (World Bank, 2022; Moers, 2015; OXFAM, 2019). Once again, this raises the question of the relationship between economic growth and income inequality. Unlike Kuznets (1955), who established an inverted U-shaped relationship between economic growth and income inequality, some authors find a linear, positive or negative relationship between the two variables. This lack of consensus makes it worthwhile to look for mechanisms that could provide a clear explanation. Among these potential factors, the environmental protection policy in a context of sustainable development deserves particular attention. Indeed, the environmental resources on which depend most of poor people in developing countries are increasingly subject to access restrictions, notably through the regulation of forest areas, fishing zones, etc. (Doumbe-Billé, 2004).

Such arrangements can exacerbate inequalities, at least in the short term, which in turn affect the quality of the environment (Boyce, 1994; Coondoo and Dinda, 2008; Dinda and Coondoo, 2006; Golley and Meng, 2012; Grunewald et al., 2011; Torras and Boyce, 1998; Wolde-Rufael and Idowu, 2017). The resulting environmental degradation has a negative impact on economic growth as

showed by several authors (Azam, 2016; Chatterjee, 2009; Etchie et al., 2017; Gallagher, 2005; Lima et al., 2015; Loizeau et al., 2018; Pao and Tsai, 2010; Reddy and Behera, 2006; Usman et al., 2019). Although an abundant literature exist on these two-way relationships, particularly between economic growth and the environment (Kozluk, T., & Zipperer, V., 2015) on the one hand, and between environmental quality and inequality (Jha et al., 2019; Boyce, 1994; Torras and Boyce, 1998; and Magnani, 2000) on the other, few authors have focused on the relationship between economic growth, income inequality and environmental quality. Those who have done so have favoured the human capital channel (Aloi and Tournemaine, 2013; Constant, 2019).

Unlike the latter, this paper identifies the role of environmental quality in the relationship between economic growth and income inequality. In particular, it analyses the effect of environmental policy on the relationship between economic growth and income inequality in sub-Saharan Africa.

The contribution of this paper is threefold. Firstly, this research makes an empirical contribution insofar as most have adopted a theoretical approach. Then, the paper focuses on the impact of environmental policy on the link between economic growth and income inequality. Finally, the paper focuses on sub-Saharan African countries characterised by weak environmental policies and high level of poverty.

The rest of the paper is structured as follows. The second section specifies the model and deal with the data. The third section presents and discusses the main results of the study. The final section draws out the policy implications and concludes the paper.

2. METHODOLOGY

2.1 Model Specification

To analyse the effect of environmental policy on the relationship between economic growth and income inequality, we use a panel threshold effect model. As a tool adapted to the analysis of non-linear phenomena (Belguith et al., 2017), this model considers context heterogeneity. The panel smooth threshold regression (PSTR) model proposed by Gonzales et al (2005) and revised in 2017 is the one used in this paper. It used for modelling situations where the transition from one regime to another is gradual.

The PSTR model takes the following form:

$$y_{it} = \mu_i + \beta_0 X_{it} + \beta_1 X_{it} G(q_{it}, \gamma, C) + e_{it} \quad (1)$$

Where y_{it} represents the dependent variable (in our case it is the Gini index), μ_i is the vector of individual fixed effects and $G(q_{it}, \gamma, C)$ denotes the transition function associated with a transition variable q_{it} , a threshold parameter C and a smoothing parameter γ . $X_{it} = (X_{it}^1, \dots, X_{it}^k)$ is the matrix of k explanatory variables (the control variables) containing no lagged endogenous variables, $\beta = (\beta_1, \dots, \beta_k)$ is the vector of coefficients and ε_{it} is the independent and identically distributed iid $(0, \sigma_i^2)$ error term. The index $i = 1 \dots N$ indicates the individual dimension corresponding to the countries and the index $t = 1 \dots T$, the temporal dimension. Theoretically, a smooth transition mechanism between regimes can be modelled using various transition functions if they are continuous and integrable on the interval $[0, 1]$.

González et al (2005) proposed a logistic transition function of order m :

$$G(q_{it}, \gamma, C) = [1 + \exp(-\gamma \prod_{j=1}^m (q_{it} - C_j))]^{-1} \epsilon [0; 1] \quad (2)$$

With $\gamma > 0, C_1 < \dots < C_m$ where $(C_1; C_2; \dots; C_m)$ is an m -dimensional vector of the threshold parameters. When $\gamma \rightarrow \infty$, the transition becomes abrupt and the PSTR model converges to a PTR with two regimes (left, respectively right of C). On the other hand, for $\gamma \rightarrow 0$, $G(\cdot) \rightarrow 1/2$, and the PSTR estimate joins that of a panel with individual fixed effects. For γ values between these two extremes, the transition is smooth between the two regimes.

The empirical specification is as follows:

$$Gini_{it} = \delta_i + \alpha_1 pole_{it} + \alpha_2 \ln pib_{it} + \alpha_3 \ln pib_{it}^2 + \alpha_4 Corrup_{it} + \alpha_5 \lnouvcom_{it} + [\alpha_6 pole_{it} + \alpha_7 \ln pib_{it} + \alpha_8 \ln pib_{it}^2 + \alpha_9 Corrup_{it} + \alpha_{10} \lnouvcom_{it}] G(q_{it}, \gamma, C) + e_{it} \quad (3)$$

With:

$Gini_{it}$: Gini inequality index for country i at period t ;

$\ln pib_{it}$: neperian logarithm of the gross domestic product per capita of country i in period t ;

$Pole_{it}$: index of the state of environmental policy in country i at period t (transition variable q_{it}).

We use the sub-indicators of the Environmental Performance Index (EPI) as environmental policy variables. These are a) the adjusted carbon dioxide performance index (IPCO2) which assesses the performance of countries in complying with policies related to the reduction of CO2 emissions ; b) the animal and plant species protection performance index (IPPEAV), which assesses countries' performance in complying with policies relating to the protection of animal and plant resources; and c) the ozone layer protection performance index (IPPCO), which indicates countries' level of performance in complying with policies relating to the protection of the ozone layer. \lnouvcom_{it} is the natural logarithm of the degree of trade openness of country i at date t . $Corrup_{it}$ is the indicator of corruption in country i at period t . This indicator belongs to the interval $[-2.5, +2.5]$. To facilitate its interpretation, we re-parameterise so that a high score indicates a strong perception of corruption. Thus, we define the variable $corrup_{it} = -WGI$ corruption indicator.

The PSTR method has several advantages. The main one being that the coefficients represent elasticities of income inequality in relation to environmental policies and economic growth and vary over time and by country. Despite its interest and flexibility, the estimation of a PSTR model requires prior tests, in particular the verification of the existence of non-linearity and the determination of the number of transitions.

2.2 Data Sources and Description

We use a panel of 26 sub-Saharan African countries covering the period from 1999 to 2021.

Table 1. Definition of variables and units of measurement

Variables	Definition	Units of measurement	Data sources
PIBH	Gross domestic product per capita	US Dollar (\$)	WDI 2023
IPPEAV	Animal and plant species protection index	Score rated on 100	EPI 2023
IPPCO	Ozone layer protection index	Score rated on 100	EPI 2023
IPCO ₂	CO ₂ emissions control index	Score rated on 100	EPI 2023
Gini	Income inequality index	Score rated on 1	WIID 2023
Corrup	Level of corruption	score belongs to the interval [-2,5; +2,5]	WGI 2023
Ouvcom	$\frac{(X+M)}{(2PIB)} * 100$ is the degree of commercial openness	%	WDI 2023

Table 2. Descriptive statistics of variables

Variables	Number of obs	Mean	Standard deviation	Minimum	Maximum
Gini	598	0.615048	0.058282	0.4945792	0.7765049
PIBH	598	1391.988	1709.417	102.598	8810.931
IPPCO	598	42.8871	12.32301	21.70335	82.66817
IPCO ₂	598	29.60408	20.77485	0	85.03919
IPPEAV	598	81.48642	12.56834	41.47988	99.63895

Source: Author, based on data WIID (2023); WDI (2023); WGI (2023) and EPI (2023)

This limited number of countries is due to the unavailability of data from other countries in the region. The data comes from the World Bank's World Development Indicator (WDI). For environmental policy, we use the Environmental Performance Index (EPI) published by the Yale Center for Environmental Law and Policy (YCELP) (Wendling et al, 2023). For the democracy indicator, the 2023 World Governance Indicator (WGI) is used. For income inequality, we use the Gini index from the 2023 World Income Inequality Database (WIID).

Table 1 describes all the variables used in the model. Descriptive statistics are presented in Table 2.

Table 2 shows that income inequality remains high in sub-Saharan Africa, with a less dispersed Gini index, averaging around 0.62 and sometimes reaching 0.78 in some countries. As for GDP per capita, it is widely dispersed around an average of US\$1,392, with some countries achieving just US\$102 billion (the minimum) and others as high as US\$8,800 billion (the maximum). The first group includes countries such as Chad, the Democratic Republic of Congo, Mali and Madagascar, with a low GDP per capita.

In terms of environmental protection, Sub-Saharan Africa has, on average, a low level of protection. However, protection of animal and plant species is high, with an average index of 81.49. In fact, animal and plant species are

better protected than atmospheric air, in particular the ozone layer with a protection index of 42.89 and CO₂ emissions with a control index of only 29.60. This performance in terms of protecting flora and fauna could be explained by several initiatives in the form of financial or technical support from international institutions and NGOs in favour of nature.

3. RESULTS AND DISCUSSION

3.1 Preliminary Tests

The application of the PSTR model requires the verification of some hypotheses based on statistical tests, in particular tests of linearity and determination of the number of regimes. Tables 3 and 4 show the results of these tests. The null hypothesis of linearity is rejected at the 1% significant level, indicating the existence of a non-linear relationship between environmental policies and income inequality. This result led to the determination of the number of regimes in the process, which is two (2), since the null hypothesis of the test could not be rejected at the 1% significant level. There is a single threshold allowing the transition from one environmental policy regime (extreme regime 1) to another (extreme regime 2). These thresholds are 42.5027, 75.9979 and 33.8038 respectively for the IPPCO, IPPEAV and IPCO₂. Table 3 shows that non-linearity holds for all transition variables.

Table 3. Non-linearity test with ippeav, ipcco and ipco2 as transition variables

Test assumptions	H0: Linear Model		
	H1: PSTR model with at least one threshold		
Transition variables	IPPCO	IPPEAV	IPCO ₂
Tests			
Wald test (LM)	W=34.83*** (0.000)	W=104.78*** (0.000)	W=53.96*** (0.000)
Fisher Tests (LMF)	F=3.475*** (0.000)	F=11.94*** (0.000)	D=5.574*** (0.000)
LRT Tests (LRT)	LRT= 35.88*** (0.000)	LRT=115.2*** (0.000)	LRT= 56.55*** (0.000)

Source: Author's calculation using Matlab2021a; ***, ** and * indicate the significance of the variables at the 1%, 5% and 10% level respectively

Since non-linearity is verified for each of the transition environmental policy variables (IPPEAV, IPPCO, IPCO₂), we will determine the number of regimes

Table 4. Number of schemes by transition variable

Test assumptions	H0: PSTR model with one regime(r=1)		
	H1: PSTR model with at least two regimes(r=2)		
transition variables	IPPCO	IPPEAV	IPCO ₂
Tests			
Wald Test (LM)	W = 27.508*** (0.002)	W=83.26*** (0.000)	W=59.31*** (0.000)
Fisher Tests (LMF)	F = 2.66*** (0.004)	F=8.93*** (0.000)	F=6.07*** (0.000)
LRT Tests (LRT)	LRT= 28.16*** (0.002)	LRT=89.66*** (0.000)	LRT= 62.37*** (0.000)

Source: Author's calculation using Matlab2021a; ***, ** and * indicate the significance of the variables at the 1%, 5% and 10% thresholds respectively

3.2 Interpretation and Discussion of the Results of the Econometric Estimates

These results show two regimes in each case. After these preliminary tests, the PSTR model is estimated, and the results are presented in Table 5.

The high values of the gamma smoothing parameters (3.8798, 8.849 and 2.3518) indicate an abrupt transition. According to Yueying (2017) the transition is abrupt and the PSTR model converges to a Hansen (1999) threshold effect model (PTR). This abrupt transition at different environmental policy thresholds indicates that SSA countries cannot be divided into an infinite number of classes, suggesting a degree of homogeneity of these countries with a similar level of inequality. It is therefore possible to clearly identify sub-groups of SSA countries on which homogenous estimates involving the three variables can be made. This result confirms the two extreme regimes. The model is globally significant and most of the variables are significant at the 1% level.

The results validate Kuznets' inverted-U relationship only in the context of strict application of environmental standards. Specifically, the results show that strict application of environmental standards reduces income inequality in sub-Saharan Africa above a wealth threshold. Indeed, the more a relatively wealthy country adopts a rigorous environmental policy, the more economic growth contributes to reduce income inequalities. In other words, economic growth can only reduce income inequality in a context of strict compliance with environmental standards. Authors such as Magnani (2000), Aloï & Tournemaine (2013) and Constant (2019) have reached the same conclusion. Magnani (2000) shows that in rich countries, economic growth improves the quality of the environment through income inequalities reduction. According to this view, a fair distribution of economic growth generated in the form of public goods (environmental quality, education, health) improves the human capital of the population, which in turn increases productivity; thereby raising the mean willingness to pay for environmental protection.

Table 5. Results for the PSTR model with ippeav, ippco and ipco2 as transition variables

Specification	Model 1		Model 2		Model 3	
TransitionVariable	IPPCO		IPPEAV		IPCO2	
Extreme regimes	r1	r2	r1	r2	r1	r2
Coef	β_0	β_1	β_0	β_1	β_0	β_1
Explanatory Variables	Gini					
IPPCO	0.019*** (2.82)	-0.02*** (-2.91)	-	-	-	-
IPPEAV	-	-	-0.002 (-1.11)	0.0034** (2.60)	-	-
IPCO ₂	-	-	-	-	0.02*** (2.82)	-0.02*** (-2.91)
Inpibh	-0.36*** (-5.85)	0.35*** (4.75)	-0.10** (-2.61)	0.04 (1.16)	-0.36*** (-5.85)	0.35*** (4.76)
Lnpibh ²	0.026*** (6.38)	-0.03*** (-5.39)	0.01** (2.73)	-0.005*** (-2.00)	0.026*** (6.38)	-0.03*** (-5.39)
Corrup	0.196*** (13.90)	-0.20*** (-14.00)	-0.013 (-1.38)	0.03*** (2.81)	0.20*** (13.91)	-0.20*** (-14.00)
Inouvcom	0.01 (0.24)	-0.02 (-0.81)	0.04*** (6.69)	-0.09*** (-10.96)	0.005 (0.24)	-0.02 (-0.81)
Smoothing Para. γ	3.88*** (0.00)		8.89*** (0.00)		2.35*** (0.00)	
Threshold Para. C	42.50*** (0.00)		75.99*** (0.00)		33.80*** (0.00)	
AIC criterion	-7.49		-7.59		-7.39	
Schwarz criterion	-7.40		-7.49		-7.39	
No observation	598		598		598	

Source: Author's calculation using Matlab2021a; ***, ** and * indicate the significance of the variables at the 1%, 5% and 10% thresholds respectively. Model1, model2 and model3 are PSTR models using respectively ippeav, ippco and ipco2 as transition variables. r1 and r2 are respectively the first and second extreme regime for each transition variable. β_0 and β_1 represent the vector of coefficients of the first and second extreme regime respectively

This idea is in line with that of Zerbo and Hien (2023), who find that economic growth will only reduce income inequalities if it greatly strengthens the capabilities of the poor.

The quality of the environment resulting from a rigorous environmental policy could constitute this channel, as a healthy environment increases life expectancy, which in turn favours investment in human capital (Harrison and Rubinfeld, 1978; Su et al., 2011; Bell et al., 2006).

This result shows that Kuznets' inverted-U relationship is not independent of environmental policy. According to the environmental Kuznets curve (EKC) highlighted by Grossman and Krueger (1995), the environment is a luxury good, as confirmed by Magnani (2000). The underlying assumption is that environmental standards become more stringent as the economy grows, due to an increasing demand for environmental quality as per capita income

rises (Antle and Heidebrink, 1995; Broad, 1994, Grunewald et al., 2017). Economic growth is a prerequisite for the application of strict environmental standards. So the socially virtuous effect (reducing inequalities) of a rigorous environmental policy only appears in economically dynamic countries. On the other hand, strict environmental standards in poor countries lead to an increase in income inequalities. Given that the poor are heavily dependent on environmental resources, any restriction on access to these resources (strict environmental regulation) leads to a deterioration in their economic conditions (Martinez-Alier, 2002; Olsson et al., 2014). Similarly, the results show that, in low-income countries, weak environmental policies are not socially inefficient because they do not increase income inequalities. This result validates the polluter's haven hypothesis, according to which developing countries adopt permissive environmental policies to attract industries

favouring job creation, which lead to income inequalities reduction.

A kind of a segmentation of environmental policy is emerging. This segmentation consists, on the one hand, in the strict application of environmental standards in developed countries (Magnani, 2000) and, on the other hand, in a permissive environmental policy in developing countries. Consequently, environmental policy design should consider a country's level of development. Any standardisation of environmental policy within a group of economically different countries would be socially inefficient because it would accentuate income inequalities.

These results contribute to the current debate on environmental policy in general, and climate policy in particular, where all the countries in the world are invited to make some commitments about greenhouse gas emissions reduction.

4. CONCLUSION

The persistence of income inequality in a context of economic growth has prompted interest in re-examining the relationship between these two phenomena, particularly in the context of sustainable development. The aim of this paper is to analyse the effect of environmental policies on the relationship between economic growth and income inequality in sub-Saharan Africa. The Panel Smooth Transition Threshold Effect (PSTR) model of Gonzalès (2005) has been used to test data from the World Bank (2023), the Yale Center for Environmental Law and Policy (YCELP) (2023), WGI (2023) and WIID (2023). Two major results were obtained. The first reveals an inverted-U relationship between economic growth and income inequality in sub-Saharan Africa when environmental standards are respected. Indeed, the more a relatively wealthy country adopts a rigorous environmental policy, the more economic growth contributes to reduce income inequalities. Thus, the socially virtuous effect of a strict environmental policy only occurs in economically dynamic countries. The second result validates the polluter haven hypothesis by indicating that weak environmental policies are not socially inefficient because they do not increase income inequalities. Consequently, the promotion of environmental policies in sub-Saharan African countries should consider the level of economic performance. Any standardisation of environmental policies on a continental or sub-

regional scale for economically different countries should be reconsidered. To ensure compatibility between economic growth and income equality, these two main results call for a segmentation of environmental policy. This involves strict application of environmental standards in developed countries and a permissive environmental policy in developing countries.

DISCLAIMER (ARTIFICIAL INTELLIGENCE)

Author(s) hereby declare that NO generative AI technologies such as Large Language Models (ChatGPT, COPILOT, etc.) and text-to-image generators have been used during the writing or editing of this manuscript.

COMPETING INTERESTS

Authors have declared that no competing interests exist.

REFERENCES

- World Bank (2022). *La Banque mondiale en Afrique*. (Visited on Septembre 30, 2023). <https://databank.worldbank.org/source/world-development-indicators>
- Moers, L. (2015). *Perspectives économiques régionales pour l'Afrique subsaharienne*. Cotonou: FMI.
- Oxfam (2019). *La crise des inégalités en Afrique de l'Ouest*. OXFORD: OXFAM.
- Kuznets, S. (1955). Economic growth and income inequality, *American Economic Review*, 45(1), 1-28.
- Doumbe-Bille, S. (2004). Les droits économiques, sociaux, environnementaux et culturels pour le développement durable. *Liaison énergie francophonie*, (64), 72-77.
- Boyce, J. K. (1994). Inequality as a cause of environmental degradation. *Ecological economics*, 11(3), 169-178. [https://doi.org/10.1016/0921-8009\(94\)90198-8](https://doi.org/10.1016/0921-8009(94)90198-8)
- Coondoo, D., & Dinda, S. (2008). Carbon dioxide emission and income: A temporal analysis of cross-country distributional patterns. *Ecological Economics*, 65(2), 375-385.
- Dinda, S., & Coondoo, D. (2006). Income and emission: a panel data-based cointegration analysis. *Ecological Economics*, 57(2), 167-181 <https://doi.org/10.1016/j.ecolecon.2005.03.028>

- Golley, J., & Meng, X. (2012). Income inequality and carbon dioxide emissions: The case of Chinese urban households. *Energy Economics*, 34(6), 1864-1872. <https://doi.org/10.1016/j.eneco.2012.07.025>
- Grunewald, N., Klasen, S., Martínez-Zarzoso, I., & Muris, C. (2011). *Income inequality and carbon emissions* (No. 92). Discussion Papers.
- Torras, M., & Boyce, J. K. (1998). Income, inequality, and pollution: a reassessment of the environmental Kuznets curve. *Ecological economics*, 25(2), 147-160. [https://doi.org/10.1016/S0921-8009\(97\)00177-8](https://doi.org/10.1016/S0921-8009(97)00177-8)
- Wolde-Rufael, Y., & Idowu, S. (2017). Income distribution and CO2 emission: A comparative analysis for China and India. *Renewable and Sustainable Energy Reviews*, 74, 1336-1345. <https://doi.org/10.1016/j.rser.2016.11.149>
- Azam, M. (2016). Does environmental degradation shackle economic growth? A panel data investigation on 11 Asian countries. *Renewable and Sustainable Energy Reviews*, 65, 175-182.
- Chatterjee, R. (2009). Economic damages from nutrient pollution create a "toxic debt".
- Etchie, T. O., Sivanesan, S., Adewuyi, G. O., Krishnamurthi, K., Rao, P. S., Etchie, A. T., ... & Smith, K. R. (2017). The health burden and economic costs averted by ambient PM2.5 pollution reductions in Nagpur, India. *Environment international*, 102, 145-156. <https://doi.org/10.1016/j.envint.2017.02.010>
- Gallagher, K. P. (2005). International trade and air pollution: Estimating the economic costs of air emissions from waterborne commerce vessels in the United States. *Journal of environmental management*, 77(2), 99-103. <https://doi.org/10.1016/j.jenvman.2005.02.012>
- Lopez, R. (2017). The environment as a factor of production: The effects of economic growth and trade liberalization 1. In *International trade and the environment* (pp. 239-260). Routledge.. <https://doi.org/10.1006/jjem.1994.1032>
- Lima, M. L., Romanelli, A., & Massone, H. E. (2015). Assessing groundwater pollution hazard changes under different socio-economic and environmental scenarios in an agricultural watershed. *Science of the Total Environment*, 530, 333-346. <https://doi.org/10.1016/j.scitotenv.2015.05.026>
- Loizeau, M., Buteau, S., Chaix, B., McElroy, S., Counil, E., & Benmarhnia, T. (2018). Does the air pollution model influence the evidence of socio-economic disparities in exposure and susceptibility? *Environmental Research*, 167, 650-661.
- Pao, H. T., & Tsai, C. M. (2010). CO2 emissions, energy consumption and economic growth in BRIC countries. *Energy policy*, 38(12), 7850-7860. <http://dx.doi.org/10.1016/j.enpol.2010.08.045>
- Reddy, V. R., & Behera, B. (2006). Impact of water pollution on rural communities: An economic analysis. *Ecological economics*, 58(3), 520-537. <https://doi.org/10.1016/j.ecolecon.2005.07.025>
- Usman, M., Ma, Z., Wasif Zafar, M., Haseeb, A., & Ashraf, R. U. (2019). Are air pollution, economic and non-economic factors associated with per capita health expenditures? Evidence from emerging economies. *International journal of environmental research and public health*, 16(11), 1967.
- Kozluk, T., & Zipperer, V. (2015). Environmental policies and productivity growth: a critical review of empirical findings. *OECD Journal: Economic Studies*, 2014(1), 155-185.
- Jha, A., Matthews, P. H., & Muller, N. Z. (2019, May). Does environmental policy affect income inequality? Evidence from the Clean Air Act. In *AEA Papers and Proceedings* (Vol. 109, pp. 271-276). 2014 Broadway, Suite 305, Nashville, TN 37203: American Economic Association.
- Magnani, E. (2000). The Environmental Kuznets Curve, environmental protection policy and income distribution. *Ecological economics*, 32(3), 431-443.
- Aloi, M., & Tournemaine, F. (2013). Inequality, growth, and environmental quality trade-offs in a model with human capital accumulation. *Canadian Journal of Economics/Revue canadienne d'économie*, 46(3), 1123-1155.
- Constant, K. (2019). Environmental policy and human capital inequality: A matter of life and death. *Journal of Environmental Economics and Management*, 97, 134-157.

- Belguith, S. O., Gabsi, M. C., & Badr, F. (2017). Effets non linéaires de la dette publique sur la croissance économique des pays MENA : Evaluation empirique à l'aide d'un modèle PSTR. Dans BCEAO, & B. C. (BCEAO) (Éd.), *Revue Monétaire* (pp. 7-24). DAKAR: BCEAO.
- Gonzalez, A., Teräsvirta, T., & Dijk, D. V. (2005). Panel Smooth Transition Regression Models. *University of Technology Sydney*, pp. 1-35.
- Gonzalez, A., Teräsvirta, T., Van Dijk, D., & Yang, Y. (2017). *Panel smooth transition regression models*.
- Yueying, L. (2017). Top Management Team Heterogeneity, Shareholding Proportion of the Largest Shareholder and Innovation Performance: A Study Based on a Panel Smooth Transition Regression Model. *Open Journal of Business and Management*, 6(1), 60-76.
- Hansen, B.E. (1999), Sample splitting and threshold estimation, *Econometrica*, in press.
- Zerbo, A., & Hien, L. (2023). *Impact de la croissance économique sur les inégalités de revenus: un modèle d'analyse*. Working Paper, DT/21/2023. Innove Center.
- Harrison Jr, D., & Rubinfeld, D. L. (1978). The distribution of benefits from improvements in urban air quality. *Journal of Environmental Economics and Management*, 5(4), 313-332.
- Su, J. G., Jerrett, M., De Nazelle, A., & Wolch, J. (2011). Does exposure to air pollution in urban parks have socioeconomic, racial or ethnic gradients ? *Environmental Research*, 111(3), 319-328.
- Bell, M. L., Davis, D. L., Gouveia, N., Borja-Aburto, V. H., & Cifuentes, L. A. (2006). The avoidable health effects of air pollution in three Latin American cities: Santiago, Sao Paulo, and Mexico City. *Environmental research*, 100(3), 431-440.
- Grossman, G. M., & Krueger, A. B. (1995). Economic growth and the environment. *The quarterly journal of economics*, 110(2), 353-377.
- Antle, J. M., & Heidebrink, G. (1995). Environment and development: theory and international evidence. *Economic Development and Cultural Change*, 43(3), 603-625.
- Broad, R. (1994). The poor and the environment: friends or foes?. *World development*, 22(6), 811-822.
- Grunewald, N., Klasen, S., Martínez-Zarzoso, I., & Muris, C. (2017). The trade-off between income inequality and carbon dioxide emissions. *Ecological Economics*, 142, 249-256.
- Martinez-Alier, J. (2002). *The Environmentalism of the poor: a study of ecological conflicts and valuation*. Edward Elgar Publishing.
- Olsson, L., Opondo, M., Tschakert, P., Agrawal, A., Eriksen, S.H., Ma, S., Perch, L.N., Zakieldean, S.A., 2014. Livelihoods and poverty. In: Field, C.B., Barros, V.R., Dokken, D.J., Mach, K.J., Mastrandrea, M.D., Bilir, T.E., Chatterjee, M., Ebi, K.L., Estrada, Y.O., Genova, R.C., Girma, B., Kissel, E.S., Levy, A.N., MacCracken, S., Mastrandrea, P.R., White, L.L. (Eds.), *Climate Change 2014: Impacts, Adaptation, and Vulnerability. Part A: Global and Sectoral Aspects. Contribution of Working Group II to the Fifth Assessment Report of the Intergovernmental Panel on Climate Change*. Cambridge University Press, Cambridge, United Kingdom and New York, NY, USA, pp. 793–832.

Disclaimer/Publisher's Note: The statements, opinions and data contained in all publications are solely those of the individual author(s) and contributor(s) and not of the publisher and/or the editor(s). This publisher and/or the editor(s) disclaim responsibility for any injury to people or property resulting from any ideas, methods, instructions or products referred to in the content.

© Copyright (2024): Author(s). The licensee is the journal publisher. This is an Open Access article distributed under the terms of the Creative Commons Attribution License (<http://creativecommons.org/licenses/by/4.0>), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Peer-review history:

The peer review history for this paper can be accessed here:
<https://www.sdiarticle5.com/review-history/127028>